



The transmission effects of the subprime crisis to emerging markets: A global VAR analysis

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Abstract

The objective of this paper is to analyze the transmission cycle in the United States in emerging markets. In this work we present a quarterly global model (GVAR) that combines individual economies vector error-correcting models in which the domestic variables are related to the country-specific foreign variables. The global VAR (GVAR) model is estimated for 32 countries over the period 1980-2013. It has the advantage of studying the effect of shocks from the U.S to emerging markets in particular the interdependence between national and international cycles. In addition to generalized impulse responses, the current paper considers the use of the GVAR for structural impulse response analysis with focus on external shocks for the emerging economy, particularly in response to shocks to the U.S. The results confirm that the U.S plays an important role in the transmission of business cycles and an economic recession affecting the U.S tends to affect emerging economies.

Keywords: Subprime crisis, GVAR, generalized impulse responses, emerging markets

Introduction

The process of globalization has been particularly significant acceleration in emerging from the 90s (Loots, 2002). An expected effect was observed an increase in the degree of synchronization of the cycle in these countries with the rest of the world (mainly the major developed countries).

Emerging countries would thus become sandstone globalization more vulnerable to the volatility of the international economy. However, as shown by the work of Kose *et al.* (2003), Akin and Kose (2008), Crucini *et al.* (2008) and Bildirici (2010) far from leading to the formation of a global cycle, globalization seems to have accompanied regional cycles increasingly marked.

The theory of decoupling economic conditions in emerging countries vis-à-vis that of the developed countries is only a manifestation of this result observed before the global financial crisis due to the collapse of the U.S market for subprime mortgages.

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Various factors have been put forward to explain this evolution. First, many developing countries have experienced rapid development since the 90s. This has contributed, on the one hand, to strengthen their resilience and, secondly, to intensify their mutual relations (through the growth of commercial exchange, called - South-South).

Second, these countries have adopted macroeconomic frameworks with the objectives of stabilizing inflation and control public debt levels.

The objective of this paper is to study the interdependence between the United States and emerging markets. We try to answer the question whether the global economic cycles are converging, or instead whether emerging markets have managed to decouple business cycle fluctuations in the U.S More specifically, we empirically quantify the contemporary impacts of external shocks and to predict future effects of shocks from the U.S to emerging markets.

Thus, we hope to achieve two goals in this work: one is to contribute to the economic literature on the topic; and the other is to have a more reliable yardstick available in order to explore the effects of the subprime crisis to emerging markets by using a Global VAR model during the last period.

The outline of the remainder of the paper is as follows. Section two presents overview of literature. Section three presents the model of the economy and gives the numerical methods used to explore the implications for empirical research. Section four summarizes the main empirical results and finally we present the discussion and conclusion. The paper has additional tables in the appendices.

Review of literature

Several empirical studies have tried to address this problem. These studies are based on simple methods that involve measuring the correlation between these countries and on more sophisticated models using VAR models and factor models.

The first works were based on cross-correlations of the growth rates of activity. Helbling and Bayoumi (2003) find that the correlation coefficients between the United States and the G-7 countries have decrease during the period 1973 to 2001. Similarly, Heathcote and Perri (2004) show that the correlations of output, consumption and investment between the United States and Europe, Canada and Japan are lower during 1986-2000 than in 1972-1985.

Other empirical studies by Frankel and Rose (1998), using panel data confirm the existence of a positive relationship between the intensification of bilateral trade and the correlation of the business cycle, which justifies the role of integration in the synchronization of the business cycle in developed countries. In this regard, Clark and Wincoop (2001), Otto *et al.* (2001), Calderon *et al.* (2002), Baxter and Kouparitsas (2004) argue that trade is the main factor that determines the degree of synchronization of the business cycle (see figure 1 and 2 in appendix).

To explain the cause of stability correlations global integration, Stock and Watson (2005) use the FSVAR model (factor augmented vector auto-regressions) to distinguish international common shocks and idiosyncratic shocks. They confirm that the decrease in the synchronization of business cycles is due to the regression of common shocks on the global economy. Canova (2003) studies the transmission of shocks from the United States on Latin American economies. Based on the VAR model with restrictions on the signs, he finds that the significant effect between the two economies is explained by financial transmission channels. In G-7 countries, they find that output

growth is influenced by common shocks and in particular by the significant fluctuations in the price of oil.

In other studies the carried out by the IMF (2007), using panel regressions, the SVAR model and the dynamic factors in a sample of 130 developed and developing countries, shows that despite the synchronization between cycles decreased over time, the role of the United States in the transmission cycles of other countries is especially important for neighboring countries. It also shows the importance of regional shocks by comparing the domestic factors and external shocks in volatility cycles. Using the same model Kose *et al.* (2008) analyzes the evolution of the degree of cyclical inter-dependence; they show the decrease of the influence of common factors to regional factors.

In general, most recent studies based on the dynamic factor model and VAR models to examine the relationships between countries. But the inconvenience of these models is that the former does not allow us to study several transmission channels, while the latter face a problem of identifying factors level.

The model structure and estimations

Patterns for co-movements, for the study of economic cycles between countries, have become more pronounced over the past two decades owing to increased economic and financial integration, with important implications for macroeconomic policy spillovers across countries.

To investigate the mechanisms of transmission of shocks globally, we propose a relatively new approach based on global VAR (GVAR) proposed by Pesaran *et al.* (2004) and developed by Dees *et al.* (2007). The originality of this model is that it takes account of global interdependence in a transparent manner by integrating a large number of countries. It also takes into account of the international transmission mechanisms channels.

Structural of global VAR (GVAR) framework

In practice, it is assumed that there are $N + 1$ countries (regions) in the global economy, indexed by $i = 0, 1, 2, \dots, N$. « Country 0 » is adopted as the reference country (the United States in this study). For each country (region), it is assumed that the country-specific variables are connected by the global economic variables by VARX*(p_i, q_i) model as follows:

$$\Phi_i(L, p_i)x_{it} = a_{i0} + a_{i1}t + \gamma_i(L, q_i)d_t + \Lambda_i(L, q_i)x_{it}^* + u_{it} \dots \dots \dots (1)$$

With x_{it} the vector of variables to model dimension $k_i \times 1$; d_t vector common to all countries like for example the international oil price variables; x_{it}^* the vector of foreign variables specific to country i of dimension $k_i^* \times 1$; $\Phi_i(L, p_i)$ and $\Lambda_i(L, q_i)$ are polynomial matrices of dimensions $k_i \times k_i$ and $k_i \times k_i^*$ with the lag (L) and represent the coefficients of the variables in the domestic and foreign countries respectively; a_{i0} and a_{i1} are vectors of dimension $k_i \times 1$ coefficients of variables trend. $\gamma_i(L, q_i)$, dimension $k_i \times k_d$, matrix polynomial coefficients of international variables d_t ; u_{it} of dimension $k_i \times 1$, the vector of idiosyncratic shocks specific to the country. It is assumed that the idiosyncratic shocks u_{it} are uncorrelated with mean 0 and matrix non-singular covariance:

Specific foreign country variables $u_{it} \approx iid(0, \sum_{ii})$ x_{it}^* are constructed from weighted averages of the corresponding relative to other countries variables. They measure the impact of trade partners on the economy in question. These variables are obtained as follows:

$$x_{it}^* = \sum_{j=0}^N \omega_{ij} x_{jt},$$

when $\omega_{ii} = 0$. ω_{ij} represents the share of country j in the trade (Exports + Imports/ Total of trade) of country i.

After selecting the lag p_i and q_i in each country by the AIC (assuming 2 as maximum lag), we estimate the models VARX* separately for each country for the possibility of cointegration between x_{it} , x_{it}^* and d_t . Once the country-specific models estimated, all endogenous variables in

the global economy are collected together in a vector $x_t = (x'_{0t}, x'_{1t}, \dots, x'_{Nt})$ with: $k = \sum_{i=0}^N k_i$ that represents the number of variables in the endogenous global model. For this, the model VARX* can be rewritten as follows:

$$A_i(L, p_i, q_i) z_{it} = \varphi_{it} \quad \text{for } (i=0, 1, 2, \dots, N) \quad \dots \dots \dots (2)$$

Where $A_i(L, p_i, q_i) = [\phi_i(L, p_i), -\Lambda_i(L, q_i)]$, $z_{it} = \begin{pmatrix} x_{it} \\ x_{it}^* \end{pmatrix}$,

$$\varphi_{it} = a_{i0} + a_{i1}t + \gamma_i(L, q_i)d_t + u_{it}$$

Let $p = \max(p_0, p_1, \dots, p_N, q_0, q_1, \dots, q_N)$, and construct $A_i(L, p)$ from $A_i(L, p_i, q_i)$. Also note that:

$$z_{it} = W_i x_t, \quad i=0, 1, 2, \dots, N \quad \dots \dots \dots (3)$$

With W_i is a $(k_i + k_i^*) \times k$ matrix defined by the country specific weights W_{ij} . With the above notations (2) can be written equivalently as:

$$A_i(L, p) W_i x_t = \varphi_{it}, \quad i=0, 1, \dots, N,$$

and then stacked to yield the VAR(p) model in x_t :
The GVAR is as follows:

$$G(L, p)x_t = \varphi_t \quad \dots \dots \dots (4)$$

Specification and estimation of the country-specific models

Dataset

The proposed model takes into account 32 countries, including 28 countries, the US, China, Japan and the UK and 28 countries aggregated into regions covering the period 1980-2013 with quarterly data (see table A in appendix). The GVAR model includes the following variables: the real output², the rate of inflation³, the exchange rate in terms of US dollars⁴, real equity prices⁵, the short interest rates⁶, the long interest rates⁷ and oil prices.

² $y_{it} = \ln(\text{GDP}_{it}/\text{CPI}_{it})$
³ $\pi_{it} = p_{it} - p_{it-1}$ with $p_{it} = \ln(\text{CPI}_{it})$
⁴ $\text{Ln}(e_{it}) - p_{it}$
⁵ $q_{it} = \text{Ln}(\text{EQ}_{it}/\text{CPI})$
⁶ $\rho_{it}^S = 0.25 * \ln(1 + R_{it}^S/100)$
⁷ $\rho_{it}^L = 0.25 * \ln(1 + R_{it}^L/100)$

These data are all collected on the website of the International Monetary Fund and from International Financial Statistics. For countries where GDP is available only annually, we proceeded to the method of interpolation used by Dees *et al.* (2007), we estimated the GVAR model using the Gauss Program.

Estimation and GVAR model specification

To estimate the country specific models we have carried out preliminary tests such as tests of stationarity (ADF test and test WS) and the exogeneity test (Johansen, 1992 and Harbo *et al.*, 1998). We noticed that introduced endogenous and exogenous variables are integrated of order 1. The assumption of weak exogeneity is not rejected for most variables confirming the hypothesis of weak exogéinité. Indeed, we found only 7 cases out of 159 cases are significant at the 5% level⁸.

As we said the key to our strategy is the weak exogeneity assumption of x^*_{it} . This test involves testing the significance of estimating the error correction term in the auxiliary equation of variables specific of foreign countries, x^*_{it} . In particular, each 1th element of x^*_{it} the following regression is carried out:

$$\Delta x^*_{it,l} = \mu_{il} + \sum_{j=1}^{r_i} \gamma_{ij,l} ECM^j_{i,t-1} + \sum_{k=1}^{s_i} \varphi_{ik,l} \Delta x_{i,t-k} + \sum_{m=1}^{n_i} \varrho_{im,t} \Delta \tilde{x}^*_{i,t-m} + \varepsilon_{it,l} \dots\dots (5)$$

Where $ECM^j_{i,t-1}$, $j = 1, 2, \dots, r_i$ are the estimated error correction terms corresponding to the r_i cointegrating relations found for the i^{th} country model and $\Delta \tilde{x}^*_{i,t-m} = (\Delta x^*_{it}, \Delta(e^*_{it} - p^*_{it}), \Delta p^0_{it})'$. The test for weak exogeneity is an F-test of the joint hypothesis that $\gamma_{ij} = 0, j = 1, 2, \dots, r_i$ in the above regression. The lag orders s_i and n_i , need not be the same as the orders p_i and q_i of the underlying country-specific VARX* models. In our case, for this test, we imposed the order of 2 lag on all foreign variables and we assumed that $s_i = p_i$.

The dynamic analysis of the GVAR model: impulse response functions

One of the important tools in the analysis of dynamical systems is the impulse response function that characterizes the possible reaction of the system to different future periods to the effect of shock variables in the model. To do so we used the general impulse response function developed by Koop *et al.* (1996) in their non-linear model. It allows studying the dynamics of transmission of shocks. It considers the impact associated with the error of a single variable on the variable 1th of i^{th} model using their historical distributions observed. It is defined as follows:

$$GI_{x,\varepsilon_{il}}(n, \sqrt{\sigma_{ii,ll}}, I_{t-1}) = E(x_{t+n/\varepsilon_{il}t = \sqrt{\sigma_{ii,ll}}, I_{t-1}}) - E(x_{t+n/I_{t-1}}) \dots\dots\dots(6)$$

With I_{t-1} shows all available information (t-1) and $\sigma_{ii,ll}$ designates the variance of the error term $\varepsilon_{ii,ll}$. This function allows studying the dynamics of the global model and estimating the magnitude of the effects of shocks to foreign variables: negative shock in equity prices in the U.S, a positive shock to oil prices and negative demand shock to U.S. In our study we limit the simulations to 2008.

Empirical results

Figure 3 shows the response of macroeconomic variables following a negative shock to stock prices in the United States over a period of two years using the bootstrap technique. This shock is equivalent to a 5% decrease in share price. It is immediately sent to different markets. This

⁸ Test results are available on request from the authors

confirms the important role of the market capitalization of the U.S in the global economy and the importance of the financial channel in the transmission of shocks. This effect is more pronounced in India, the region of emerging Asia and Latin America. The shock response of African financial markets is relatively low. This is an expected result if we take into account the low level of international financial integration in the region. Then, a decline in GDP of 0.8% U.S shows the dependence of the actual activity of these regions of U.S external funding. Indeed, sudden stops in these regions are the main cause of the decline in GDP. This can be explained by the withdrawal of foreign investors from their financial investments.

Shock associated with a decline in demand from the U.S (figure 4) show that this shock is equivalent to a decrease of 0.4 % of GDP in the US. This will negatively impact but slowly transmitted to the GDP of Latin America and emerging Asia. For Africa, we have seen a decline in GDP more important than the shock associated with the equity prices. This result suggests that the trade channel between this region and the United States is more important than the financial channel. The decrease in GDP in the U.S will affect demand for raw materials as well as their prices.

China and India are not affected by the shock. This can be explained by the fact that these economies have become more independent of the U.S, boosting intra-regional trade. This idea was confirmed by IMF (2007).

Figure 5 show the oil price shock. The 12 % increase in oil prices has a negative effect on the GDP of the U.S decreases by 0.08 % after one quarter before leveling quickly. The impact on oil prices negatively affects all emerging markets. This suggests that the dependence of these countries on oil is needed to manufacture all kinds of products. In our sample, Saudi Arabia and Mexico are the two largest exporters of oil, they recorded low growth, but the effect wears off quickly. Concerning equity prices, as expected, the reaction will be negative in all regions.

Discussion and conclusion

This paper updates and extends the GVAR model of Dees *et al.* (2007). The subprime crisis has given rise to an important debate on the extent of decoupling of the state of emerging markets compared to that of developed countries, especially with the cycle of the United States. In this work, our aim was to show how disturbances from the U.S are transmitted to other countries of the world and especially in emerging markets. We used a new approach, the GVAR model to estimate the effect of external shocks from the U.S on cycles in emerging markets. The originality of this model is to study the interdependence between domestic factors and international factors linking each country with the rest of the global economy with the trade weights, which can generate forecasts for the global economy. In this regard, we have identified the effect of two types of shocks from the U.S to emerging markets: the financial impact and real impact. We showed that the U.S has an important role in the transmission cycle in emerging markets. Indeed, whatever the nature of the shock from the E.U, they have adverse effects and persistent on cycles in emerging markets. This effect is explained by the fact that the share of U.S trade in most emerging markets is high, and even if this is not the case, these countries will be influenced by the effect of third-party market. However, the magnitude of this effect varies by country and region. It depends on the degree of integration of these countries into the world economy. Latin America, emerging Asian countries is the most affected areas. This is logical given that these regions have a direct commercial and financial link with the U.S. Comparing these two shocks, it seems clear that financial shocks are transmitted more quickly and have more serious consequences on macroeconomic variables in emerging markets except Africa. The latter is characterized to a less developed capital market today and less open. Therefore, the financial channel is more important

than the trade channel in the transmission of shocks to emerging markets. These countries remain until today heavily dependent on foreign capital flows in the form of foreign direct investment. The effect of real shock decreased since these countries were able to develop intra-regional trade and exploit other markets.

In this respect, our study was not limited to these two types of shocks. In this framework we have integrated our analysis the effect of the common stock through the variable oil prices. It negatively affects all emerging markets. This situation is exacerbated by inflationary pressures showing the dependence of these economies to oil prices.

Finally, our work has shown shortcomings. The choice of macroeconomic variables is limited. In our study, promising new tracks can be addressed in order to further expand the topics discussed in this work. We propose two possible extensions.

The first is to further research on fiscal and monetary policy in emerging markets during the effects of the subprime crisis. We can focus our research by studying the question of the effectiveness of these policies in major emerging markets. This allows us to draw the economic and financial characteristics relating to each emerging countries and compare them. Second is to develop the econometric method. Binder and Gross (2013) develop a Regime-Switching Global Vector Autoregressive (RS-GVAR) model. The RS-GVAR model allows for recurring or non-recurring structural changes in all or a subset of countries. It can be used to generate regime-dependent impulse response functions which are conditional upon a regime-constellation across countries.

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Appendix

Table A: Countries and regions in the GVAR model

United States	European area	Latin America
	Germany	Argentina
China	France	Brazil
Japan	Italy	Chile
United Kingdom	Belgium	Mexico
	Finland	Peru
	Netherlands	Colombia
Developed countries	Emerging Asian countries	Emerging Europe
	Korea	
Canada	Singapore	Poland
New Zealand	Thailand	Turkey
Norway	Philippines	
	Indonesia	
Africa	Rest of the world	
Tunisia		
Egypt	India	
Morocco	Saudi Arabia	
South Africa		

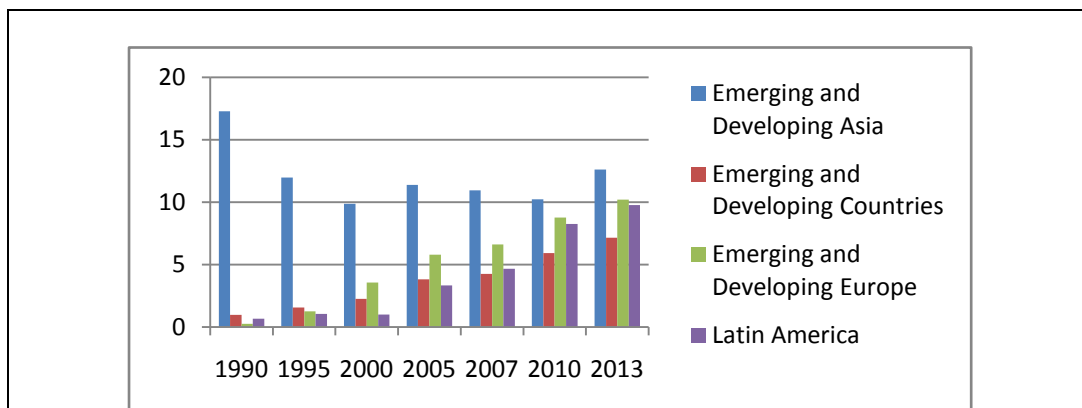


Figure 1: Export to china as a percentage of total exports

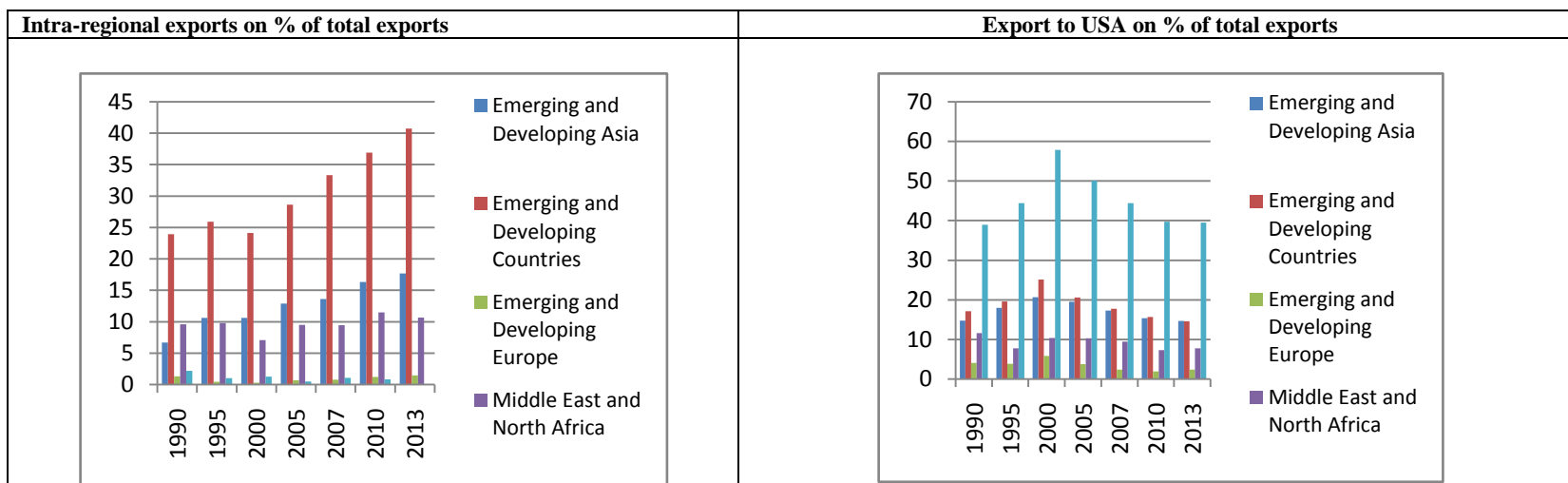


Figure 2: The share of exports across regions

Source: Autor's calculations based on Direction of Trade Statistics and IMF

Source: Autor's calculations based on Coordinated Portfolio Investment Survey and IMF

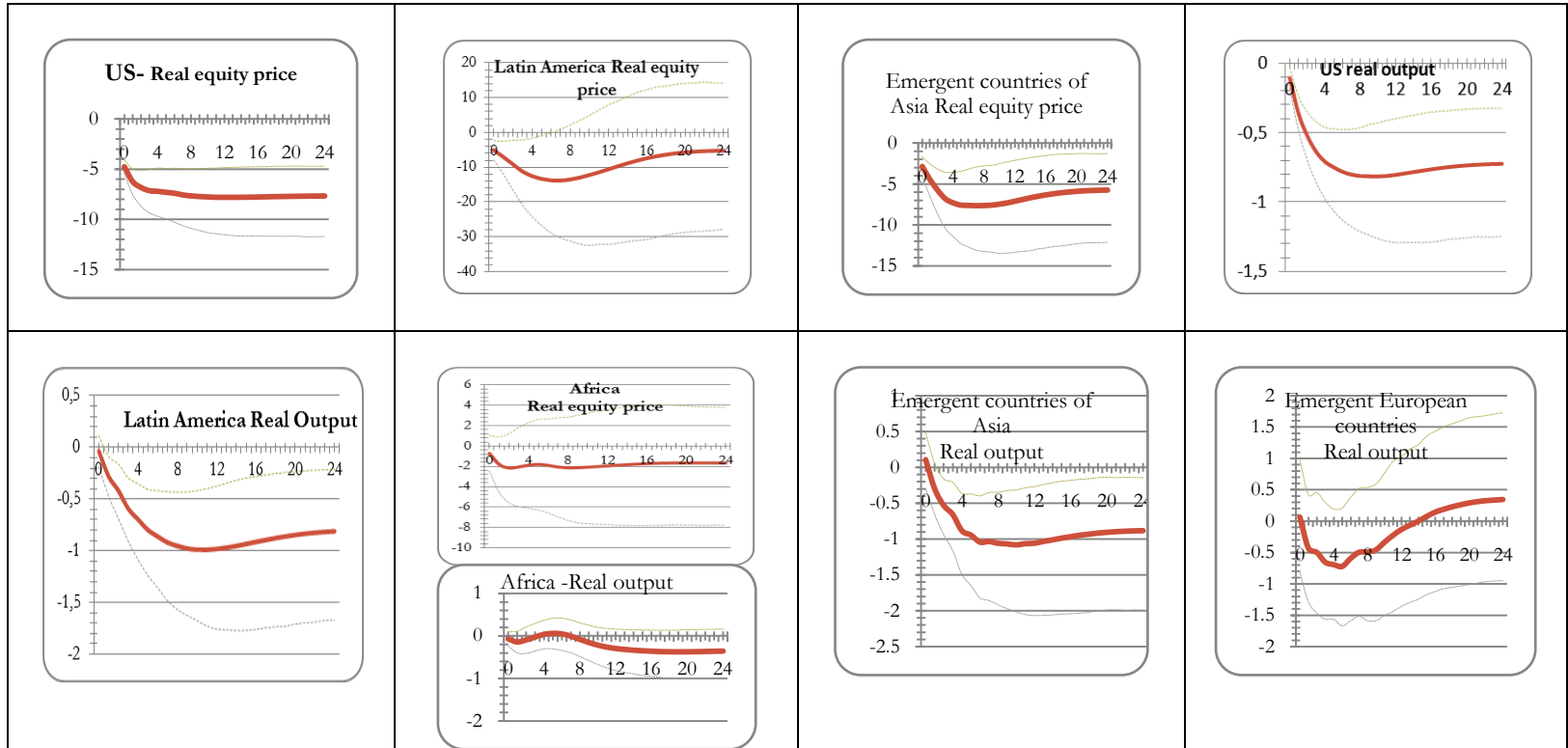


Figure 3: Generalized impulse responses of a negative unit shock to U.S. real equity prices (bootstrap mean estimates with 90% bootstrap bounds)

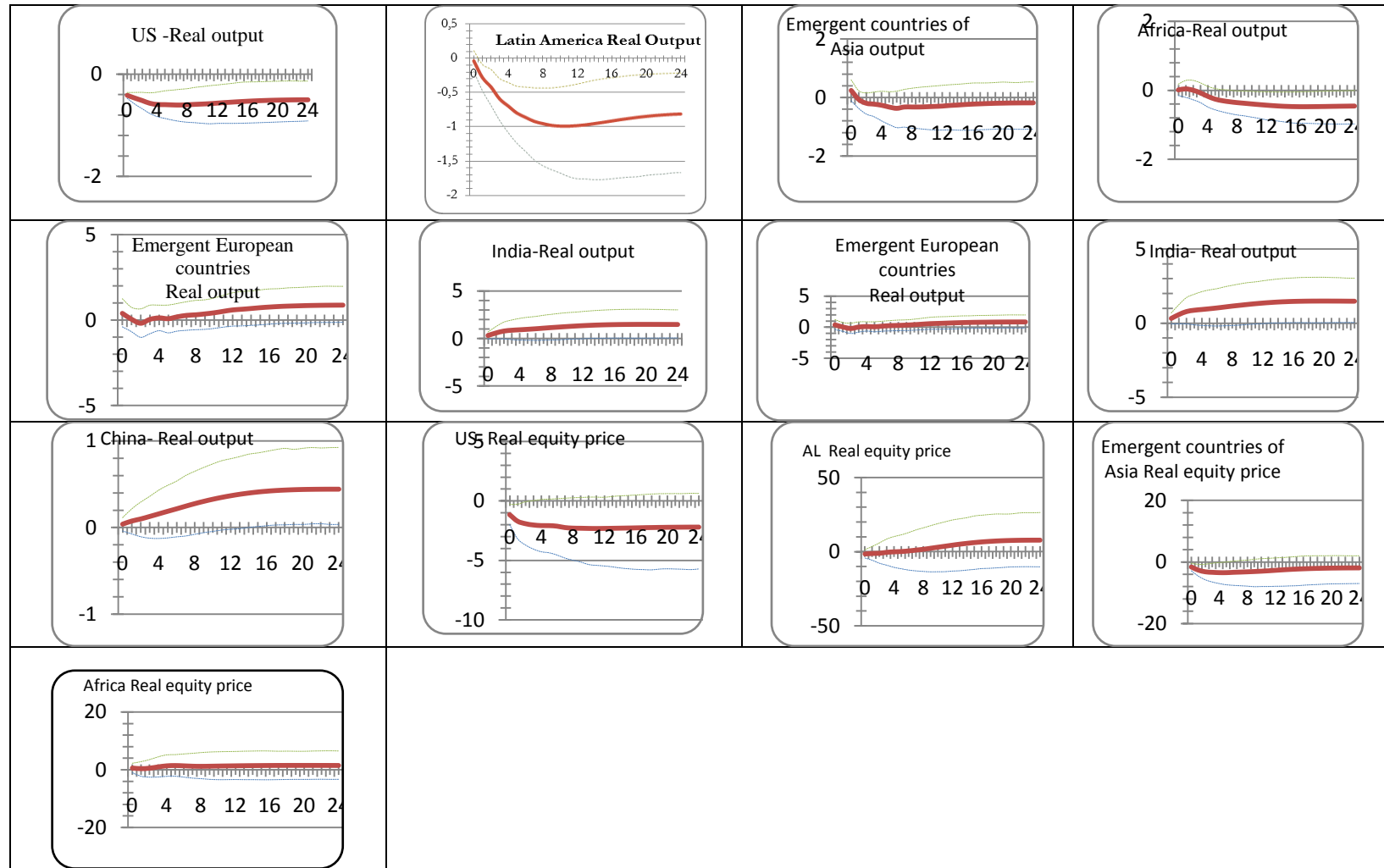


Figure 4: Generalized impulse responses of a negative shock of the demand in united state (bootstrap mean estimates)

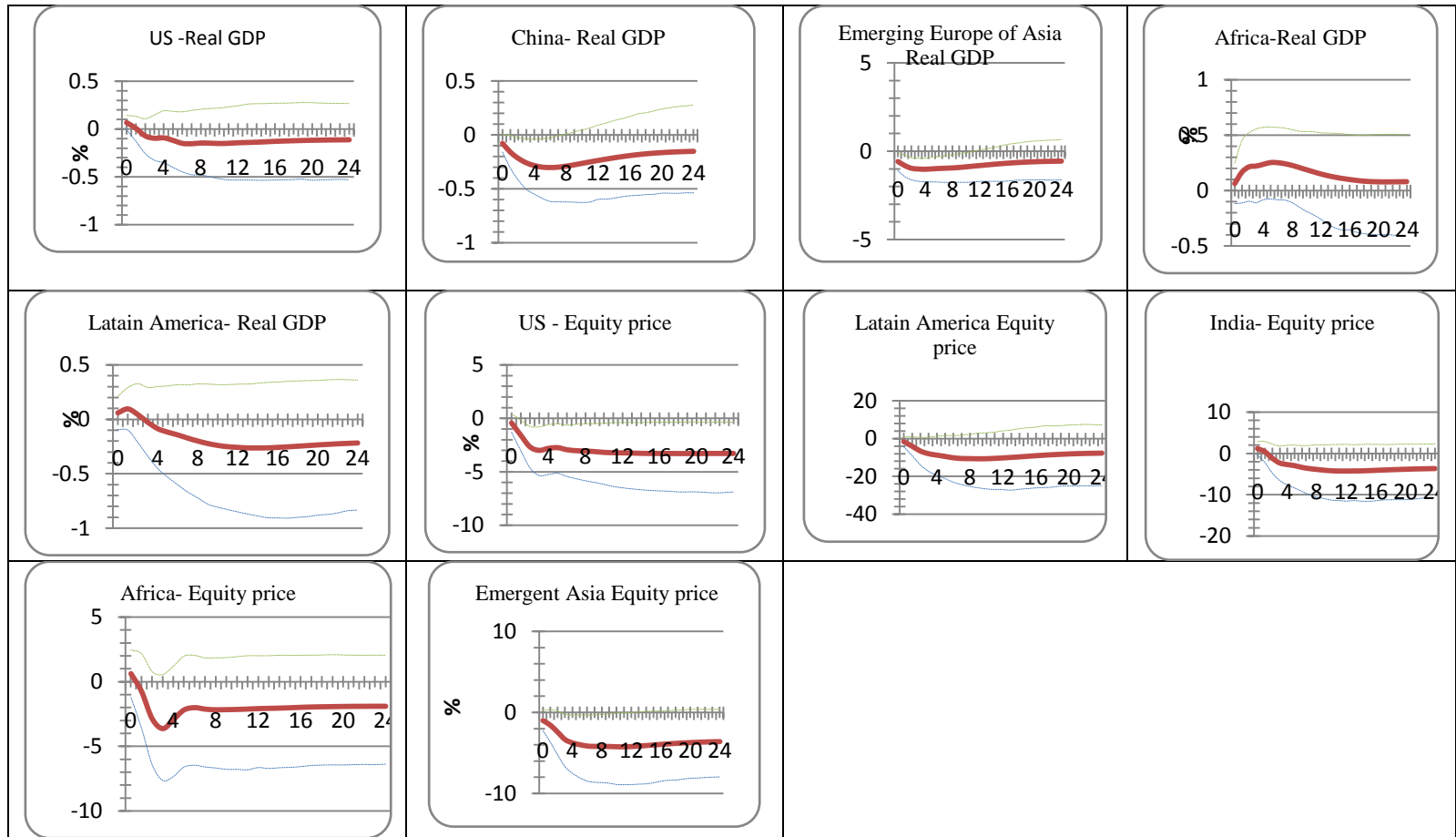


Figure 5: Generalized impulse responses of a negative shock of oil prices in united state (bootstrap mean estimates)